



The Evolution of Commodity Trios Prices and Causality Equation: In Structural Break Perspective

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ABSTRACT

This paper aims to examine the evolution of the commodity trios prices and relation among in structural break perspective for the period of 1991-2020. We used factor analysis, linearity, structural break, unit-root, cointegration tests and causality method were applied. Bai-Perron (2003) test results show industrial materials, energy and food price have one break at 2003, 2004, and 2005, respectively. To catch changing causality equation parameters clearly, it was used the moving sub-break range as 2003-2008. In post-break range, equation parameters obviously differed. In pre-break range, energy and food price have a in-directional effect on materials price, but, directional in post-break range. Lastly, in post-break range, the link among commodity trios have bi-directional relation between all commodity duos unlike pre-break range.

Keywords: Commodity Price, Time Series, Factor Analysis, Cointegration, Causality

JEL Classifications: Q02; C32; C38

1. INTRODUCTION

Commodities are tradable natural sources, which one of the production factors. Industrial raw materials, energy and agricultural production constitutes the primary categories of commodity (Matos and Wagner, 1998) and also named commodity trios. Commodity price series behaviour and relation among them are becoming increasingly complex and dynamic. Commodity prices are affected by economic, financial, geopolitical, technological and climatic events. Natural sources, belonged to the countries established on it, are distributed unequally geographically. Therefore, the current world economic system includes source rich and poor countries. Source rich countries, generally exporter countries, name the game and lead world economic activity by hand using some market mechanism or geopolitical action to manage maximization problem. It causes endless crises generation process as a facet circle. During the facet circle, it is important that understanding commodity prices and causal equation to manage risk and make decision for both of source rich and poor countries.

Commodity trios prices have a causal relationship based on production process and transportation cost. Energy commodities have a crucial role for industrial and agricultural production. Industrial raw materials, particularly metals depend on energy by smelting, transportation, mining, and production process. The relation between agricultural production (food) and industrial raw materials can be explained by using biofuel in transportation, mining, and the production process.

IMF Commodity Energy Price Index, IMF Commodity Industrial Inputs Price Index and FAO Food Price Index series calculated by international institutions, presented in Graph 1, are the most popular monitoring tools for the commodity trios price behaviour. Commodity trios price behaviour is shaped by supply and demand side factors/events.

Commodity prices traded within a range in 90's. Beginning of the 2000's, expansionary monetary policy implemented by central banks to fight recession contributed rising commodity prices. OPEC

Table 1: IMF structural classification of commodities and codes and definition of commodity price variables

| Commodities blocks ¹ | Commodities variables sub-blocks ¹ (Commodity Trios) | International institutions commodity price indexes ² | | The contents of international institutions commodity price indexes ³ | Proxy variables of commodity trios ⁴ | |
|---------------------------------|---|---|--|--|---|---|
| | | Code | Source and Name | Name | Code | Source and Name |
| Non-Fuel | Industrial Input Raw Materials (shortly Materials) | CIPI | IMF Commodity Industrial Inputs Index 2005=100 | Agricultural Raw Materials: Timber, Cotton, Wool, Rubber, and Hides indexes Metals: Copper, Aluminum, Iron Ore, Tin, Nickel, Zinc, Lead, and Uranium indexes Meat, Dairy, Cereals, Oils, Sugar Indexes | COPP | IMF Copper Price |
| | Food and Beverage (Food) | FPI | FAO Food Index 2014-2016=100 | | CPI | FAO Cereal Price Index |
| Fuel | Fuel (Energy) | CEI | IMF Commodity Energy Index 2005=100 | Oil (WTI, Brent), Natural Gas (Henry Hub), Coal indexes | WTI | IMF Texas Intermediate (WTI) ⁵ |

¹IMF Structure of Commodities, ²International Institutions Commodity Price Indexes used for determining variables sets from its contents, creating graphs and summary statistics in the study, ³The Contents of International Institutions Commodity Price Indexes used to determine proxy variables from its contents in the model, ⁴Proxy Variables instead of Commodity Trios determined using by Factor Analysis, the implication results can be see at Results section, ⁵US crude oil price

Table 2: Factor analysis results: Eigenvalues and factor loading for contents of industry inputs prices

| Variables | Factor1 | Factor2 | Factor3 | Factor4 | Factor | Eigen-value | Difference | Proportion | Cumulative |
|-----------|---------|---------|---------|---------|--------|-------------|------------|------------|------------|
| Copper | 0.990 | -0.070 | 0.002 | -0.081 | F1 | 7,7687 | 6,1843 | 0,6805 | 0,6805 |
| Aluminum | 0.826 | 0.446 | 0.076 | -0.079 | | | | | |
| Iron-ore | 0.839 | -0.398 | -0.007 | 0.165 | | | | | |
| Tin | 0.910 | -0.329 | -0.073 | 0.077 | | | | | |
| Nickel | 0.815 | 0.468 | 0.176 | -0.092 | | | | | |
| Zinc | 0.810 | 0.233 | -0.235 | -0.383 | | | | | |
| Lead | 0.954 | -0.081 | -0.129 | -0.051 | | | | | |
| Uranium | 0.802 | 0.322 | 0.147 | 0.034 | | | | | |
| Wool | 0.795 | -0.308 | -0.323 | 0.031 | | | | | |
| Rubber | 0.865 | -0.175 | 0.352 | 0.140 | | | | | |
| Cotton | 0.395 | -0.552 | 0.343 | -0.013 | F2 | 15.844 | 0.623 | 0.139 | 0.819 |
| Hides | 0.018 | 0.270 | 0.608 | 0.099 | F3 | 0.962 | 0.477 | 0.084 | 0.904 |
| Timber | 0.403 | 0.477 | -0.334 | 0.500 | F4 | 0.485 | 0.207 | 0.043 | 0.946 |

LR test: Independent versus saturated: Chi-square (78)=514,66 Prob>Chi-square=0.000

Table 3: Factor analysis: Factor Loadings for Food Price Index FAO (Meat, Dairy, Cereals, Oils, Sugar Price Index)

| Variables | Factor1 | Factor2 | Factor3 | Factor | Eigenvalue | Difference | Proportion | Cumulative |
|-----------|---------|---------|---------|--------|------------|------------|------------|------------|
| Meat | 0.870 | 0.150 | 0.172 | F1 | 4.181 | 4.043 | 0.973 | 0.973 |
| Diary | 0.916 | -0.202 | 0.090 | | | | | |
| Cereals | 0.982 | -0.094 | 0.010 | | | | | |
| Oils | 0.951 | -0.064 | -0.188 | | | | | |
| Sugar | 0.843 | 0.246 | -0.076 | | | | | |

LR test: independent versus saturated: Chi-square (3)=241.27 Prob>Chi-square=0.0000

behaviour cycle changed as the price-defense strategy again, after 9/11 2002 terrorist attack and 2003 Gulf US-Iraq War. Barsky and Kilian (2004) conclude that the changes of policy by OPEC are not exogenous, but respond to the state of the oil market and global economy. In the same way, Lavaller (2004) conclude that it is possible for OPEC members revenues by adopting a strategy. Energy, industrial materials and food price series were broken up in 2003, 2004, and 2005, respectively, due to 9/11 terrorist attack and 2003 US-Iraq war and shifted mean upside and more volatile. It can be characterized by some other factors as climate change, droughts, rising non-food agricultural production post-Kyoto (2006), income growth and changing trading behaviour with high frequency trading

AI algorithm. Commodity prices up trend, deteriorated with 2008 financial crisis and 2010 European Debt Crisis.

This paper examine the time series properties of commodity trios prices, selected proxies by factor analysis method. We applied Brock et al. (1996) was used to address linearity/non-linearity of variables. Bai-Perron (2003) structural break tests, Augmented Dickey-Fuller (ADF), Phillips-Perron (PP), Zivot Andrews (1992) allowing by one structural break, and Clemente-Montañés-Reyes (CMR) with two structural breaks unit-root tests and Johansen cointegration and VAR Granger causality test were applied.

Table 4: Factor analysis: Energy price index: Oil (WTI), coal, natural gas (Henry Hub)

| Variables | Factor1 | Factor2 | Factor | Eigenvalue | Difference | Proportion | Cumulative |
|-----------|---------|---------|--------|------------|------------|------------|------------|
| OIL | 0.942 | 0.056 | F1 | 1.783 | 1.411 | 0.888 | 0.888 |
| COAL | 0.842 | -0.326 | | | | | |
| NG | 0.432 | 0.512 | F2 | 0.372 | 0.519 | 0.185 | 1.073 |

LR test: independent versus saturated: Chi-square (3) = 48.75 Prob>Chi-square=0.0000

Table 5: Unit root test results

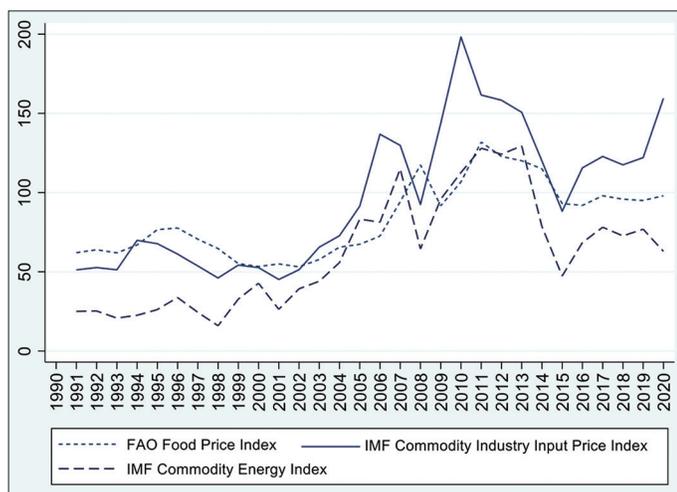
| | Dickey-fuller test | | Phillips-perron test | | Zivot andrews unit-root test with one structural break | | | Clemente-montañés-reyes unit-root test with two structural break | | | | |
|-------|--------------------|----------------|----------------------|------------|--|-------------|--------------------|--|------------------|------------|----------------------|--|
| | Trend | | Trend | | Trend | Break point | Constant and trend | Break point | Additive outlier | | Innovational outlier | |
| | Min t | Opt. Break | Min t | Opt. Break | Min t | Opt. Break | Min t | Opt. Break | Min t | Opt. Break | | |
| Level | | | | | | | | | | | | |
| COPP | -1.290 (0.633) | -1.126 (0.704) | -2.830 | 2011 | -3.994 | 2005 | 3.547* (0.002) | 2007 2010 | 5.021* (0.000) | 2004 2011 | | |
| WTI | -1.624 (0.470) | -1.533 (0.517) | -3.570 | 2011 | -4.413 | 2014 | 10.281* (0.000) | 2003 2014 | 4.910* (0.000) | 2003 2012 | | |
| CPI | -1.564 (0.501) | -1.537 (0.515) | -2.480 | 2012 | -4.161 | 2007 | 10.421* (0.000) | 2007 2014 | 4.744* (0.000) | 2005 2014 | | |

Dickey-Fuller, Philip-Perron unit root test: (), MacKinnon approximate P value for Z (t). Zivot and Andrews unit root test: Critical values for both 1%: -5.57 5%: -5.08 10%: -4.82, and trend 1%: -4.93 5%: -4.42 10%: -4.11

Table 6: Structural break test in level series of commodity prices with R

| Bai-Perron (2003) | Structural break point (for one break) |
|-------------------|--|
| WTI | 2003 |
| COPP | 2004 |
| CPI | 2005 |

Graph 1: Commodity trios price indexes calculated by international institutions



Source: FAO, IMF, WorldBank

2. LITERATURE REVIEW

The relationship among energy, materials and food prices have been investigated in related literatures. Tyner (2010) searched the relationship between energy and agricultural production prices and found a low correlation before 2005, with a peak and strengthening of the correlation in mid-2008. Chen et al. (2010) studied the relationship between oil prices and grain prices and observed an impact of oil prices on grain prices. Pala (2013) studied the relation between food and energy price by Granger causality and they found the relation between crude oil and food price equation has changed after 2008 break. Pala (2018) examined energy and food price relation

in Kyoto (2006) perspective and found there is uni-directional running from food price to energy price in post-Kyoto (2006) sub-sample. Zhang and Tu (2016) examined the effects of oil price shocks on China’s metal markets. The findings revealed significant symmetric impacts of crude oil price shocks on the metal markets. Ezeaku et al. (2021) examined the volatility of commodity prices during the COVID-19 pandemic using SVAR modeling. The results showed that copper prices initially responded positively to crude oil price shocks and then exhibited a negative response thereafter. Ji and Fan (2012) examined the influence of the crude oil market on non-energy commodity markets using a bivariate EGARCH model. The results revealed there are volatility spillover effects from the crude oil market to non-energy commodity markets. Jiang et al. (2018) explored the dynamic dependence among crude oil, agricultural raw materials, and metals using the wavelet squared coherence approach. The results showed that the oil market lags behind agricultural raw material markets but leads metal markets, while metal markets change in parallel with agricultural raw material markets. Kaulu (2021) analyzed the effects of crude oil prices on copper and maize prices using VAR and VECM models. The study did not find Granger causality running from crude oil prices to copper and maize prices.

Nelson and Plosser (1982) advocated that all macroeconomic time series have a unit root. But, Perron (1989) suggested that time series have changing pattern of it permanently due to some unique economic events. Bai-Perron (2003) are defined as a single-group change point model in the sense that all the parameters subject to the structural changes have the structural shifts at the same dates. Doornik (2022) said that “all the testing papers referenced so far take a technical approach, deriving the asymptotic properties of the proposed tests.” They focus on Yang (2017) study, discussed that trend break selection in the underspecified case is not consistent. They proposes to date a broken trend in the differenced model. Doornik (2022) the underspecified model is consistent again in that case. Doornik (2022) found that Bai-Perron (2003) algorithm valid in level in model 2, and also differenced model 2 and differenced model 3.

This paper aims to fill the gap in the literature concerning the relationship between food, energy, and materials prices, analysing

data generation process, structural break and relations among commodity trio prices, in break perspective. We preferred to use newly best proxy variable generated by factor analysis method, it has never been used before related literature.

3. METHODS AND DATA

Factor analysis method is to identify a reduced number of variables from a larger set. We preferred total variance explain criteria as 60% level and entering factor loading value criteria as up to 0.5. Brocke et al. (1996) linearity test, Bai-Perron (2003) structural break test and Augmented Dickey-Fuller (ADF), Phillips-Perron (PP), Zivot Andrews (ZA) with one structural break, and Clemente-Montañés-Reyes (CMR) with two structural breaks linear unit-root tests were applied. To determine the optimal lag length for cointegration, Akaike (AIC), Hannan and Quinn (HQIC), Schwarz’s Bayesian Information Criteria (SIC), and Final Prediction Error Criteria (FPE) test was used. The linear Johansen Cointegration and VAR Granger causality test were applied.

We used structural classification of commodities sub-blocks defined by IMF and FAO. IMF structural classification of commodities, international commodity price indexes, all contents of the indexes and selected proxy variables represent at Table 1. All data cover the period of 1991-2020 and collected from IMF, FAO and NYMEX.

4. RESULTS

We applied factor analysis method on contents of international institutions commodity price indexes to select the best proxy variables. Factor analysis results are presented in Tables 2-5. Total variance explaining ratios of commodity trios factor groups, F1, have bigger than acceptancy criteria level of 60%. We selected IMF copper price, WTI US oil price, and FAO Cereal Price Index, have the highest factor loading value, as a proxy variables instead of commodity trios; materials, energy and food price, respectively.

We applied the Augmented Dickey-Fuller (ADF), Phillips-Perron (PP), Zivot Andrews (ZA) allowing by one structural break, and

Clemente-Montañés-Reyes (CMR) with two structural breaks unit-root tests for IMF Copper Index, IMF WTI price and FAO Cereal Index. The results, presented in Table 6, indicates that the null hypothesis of a unit root cannot be rejected at the level. All series have a unit root in level. But, it is usefull and necessary that to examine data in terms of possible structural break dates. ZA and CMR test results indicates that the null hypothesis of unit root can be rejected, it suggests that all series stationary with break.

We used Bai-Perron (2003) structural break tests and the results represents at Table 7. The results verified in 2003 Gulf Warbreak in WTI price. The break reflected by ones year lagged on materials and food prices, 2004 and 2005.

Table 8 presents the outcomes of the Brock et al. (1996) non-linearity test for full sample, pre and post-break sub-sample, generated by Bai and Perron (2003) structural break test. The test results indicate that the series follow non-linear process in full sample. Otherwise, when BDS test was applied for sub-samples generated by break date points, AR(1) model residuals are not correlated and the price series follow a linear process for pre and post-break period. All series follow linear process for pre-and post-break samples.

The FPE (Akaike, 1969), AIC (Akaike, 1973), HQIC (Hannan and Quinn, 1979), and SBIC (Shwartz, 1978) criteria suggest lag lengths of 2, 3, 2, and 1, respectively. If there are conflicting results, we follow AIC, suggested by Pesaran and Pesaran (1997). It was selected as 3 for the optimal number lag length.

Table 9 shows Johansen’s cointegration test results. The trace statistics evidence of the presence of at least one cointegrating vector among and long-run relationship among energy, food and materials prices proxy variables.

The results of unit-root tests show that the series are with break stationary. We generate vector autoregression model taking into account possible breaks as 2003-2004-2005, changing period. We generated sub-samples with moving break range by 2003-2008

Table 7: BDS non-linearity test for full sample and sub-break samples

| Variables | Full sample | Epsilon/dimension | 2 | 3 | 4 | 5 | Series form |
|--------------------|-------------|-------------------|-----------------|-----------------|-----------------|-----------------|-------------|
| COPP | 1991-2020 | 4301.75 | 0.098* (0.000) | 0.187* (0.000) | 0.252* (0.000) | 0.269* (0.000) | Non-linear |
| CPI | | 41.85 | 0.105* (0.000) | 0.181* (0.000) | 0.223* (0.000) | 0.239* (0.000) | Non-linear |
| WTI | | 43.88 | 0.096* (0.000) | 0.168* (0.000) | 0.218* (0.000) | 0.233* (0.000) | Non-linear |
| Sub-samples | | | | | | | |
| WTI | 1991-2003 | 6.55 | -0.006 (0.973) | -0.025 (0.3336) | -0.047 (0.0645) | -0.126* (0.000) | Linear |
| | 2004-2020 | 27.54 | 0.051* (0.000) | 0.013 (0.3692) | -0.016 (0.3073) | -0.012 (0.3995) | Linear |
| COPP | 1991-2004 | 716.45 | -0.059* (0.002) | -0.020 (0.5035) | -0.046 (0.2031) | -0.078* (0.037) | Linear |
| | 2005-2020 | 1822.48 | 0.030 (0.061) | 0.041 (0.107) | 0.045 (0.133) | 0.065 (0.0361) | Linear |
| CPI | 1991-2005 | 10.08 | 0.067 (0.009) | 0.062 (0.1542) | 0.063 (0.2593) | 0.055 (0.3798) | Linear |
| | 2006-2020 | 36.76 | 0.028 (0.056) | -0.004 (0.8598) | 0.021 (0.5184) | 0.038 (0.3195) | Linear |

*represent 5% significance levels

Table 8: Johansen linear cointegration test (in level)

| Proxy variables | Null/alternative | Trace statistics | 5% critical value |
|-------------------------|-------------------|------------------|-------------------|
| COPP, WTI CPI, Lags (3) | $r=0/r\geq 1$ | 18.833* | 29.68 |
| | $r\leq 1/r\geq 2$ | 8.326 | 15.41 |

*presents 5% statistically significancy

Table 9: Linear granger causality test for full and moving break sub-samples

| Ho hypothesis | Chi-square-statistics | P-value | Causality direction | Chi-square-statistics | P-value | Causality direction |
|--|-----------------------|-----------|---------------------|-----------------------|-----------|---------------------|
| Case: difference stationary, with difference series | | | | | | |
| Full Sample (1991-2020) | | | | | | |
| CPI does not cause COPP | 1.538 | 0.674 | - | | | |
| WTI does not cause COPP | 4.1127 | 0.250 | - | | | |
| COPP does not cause CPI | 55.552* | 0.000 | COPP→CPI | | | |
| WTI does not cause CPI | 21.879* | 0.000 | WTI→CPI | | | |
| COPP does not cause WTI | 14.832* | 0.002 | COPP→WTI | | | |
| CPI does not cause WTI | 10.418* | 0.015 | CPI→WTI | | | |
| Case: break stationary, with level series, comparing moving break date sub-samples | | | | | | |
| Break Date: 2007 | | 1991-2007 | | | 2008-2020 | |
| CPI does not cause COPP | 2.919 | 0.404 | - | 24.427* | 0.000 | CPI→COPP |
| WTI does not cause COPP | 11.921* | 0.008 | WTI→COPP | 13.262* | 0.000 | WTI→COPP |
| COPP does not cause CPI | 147.11* | 0.000 | COPP→CPI | 27.39* | 0.000 | COPP→CPI |
| WTI does not cause CPI | 67.128* | 0.000 | WTI→CPI | 25.576* | 0.000 | WTI→CPI |
| COPP does not cause WTI | 14.287* | 0.003 | COPP→WTI | 50.706* | 0.000 | COPP→WTI |
| CPI does not cause WTI | 6.4871 | 0.090 | CPI→WTI | 56.126* | 0.000 | CPI→WTI |
| Break Date: 2006 | | 1991-2006 | | | 2007-2020 | |
| CPI does not cause COPP | 7.213 | 0.065 | - | 9.733* | 0.021 | CPI→COPP |
| WTI does not cause COPP | 29.22* | 0.000 | WTI→COPP | 3.192 | 0.363 | - |
| COPP does not cause CPI | 106.76* | 0.000 | COPP→CPI | 17.275* | 0.001 | COPP→CPI |
| WTI does not cause CPI | 63.447* | 0.000 | WTI→CPI | 19.746* | 0.000 | WTI→CPI |
| COPP does not cause WTI | 28.471* | 0.000 | COPP→WTI | 40.959* | 0.000 | COPP→WTI |
| CPI does not cause WTI | 15.409* | 0.001 | CPI→WTI | 30.659* | 0.000 | CPI→WTI |
| Break Date: 2005 | | 1991-2005 | | | 2006-2020 | |
| CPI does not cause COPP | 3.115 | 0.374 | - | 9.530* | 0.023 | CPI→COPP |
| WTI does not cause COPP | 15.362* | 0.002 | WTI→COPP | 2.749 | 0.432 | - |
| COPP does not cause CPI | 364.62* | 0.000 | COPP→CPI | 30.873* | 0.000 | COPP→CPI |
| WTI does not cause CPI | 377.80* | 0.000 | WTI→CPI | 21.973* | 0.000 | WTI→CPI |
| COPP does not cause WTI | 24.936* | 0.000 | COPP→WTI | 32.613* | 0.000 | COPP→WTI |
| CPI does not cause WTI | 18.993* | 0.000 | CPI→WTI | 25.839* | 0.000 | CPI→WTI |
| Break Date: 2004 | | 1991-2004 | | | 2005-2020 | |
| CPI does not cause COPP | 3.123 | 0.373 | - | 0.818 | 0.845 | - |
| WTI does not cause COPP | 4.310 | 0.230 | - | 0.416 | 0.937 | - |
| COPP does not cause CPI | 2980.0* | 0.000 | COPP→CPI | 47.570* | 0.000 | COPP→CPI |
| WTI does not cause CPI | 1294.3* | 0.000 | WTI→CPI | 20.887* | 0.000 | WTI→CPI |
| COPP does not cause WTI | 32.186* | 0.000 | COPP→WTI | 23.446* | 0.000 | COPP→WTI |
| CPI does not cause WTI | 31.382* | 0.000 | CPI→WTI | 18.328* | 0.000 | CPI→WTI |
| Break Date: 2003 | | 1991-2003 | | | 2004-2020 | |
| CPI does not cause COPP | 21.362 | 0.523 | - | 0.625 | 0.474 | - |
| WTI does not cause COPP | 20.130 | 0.134 | - | 0.321 | 0.361 | - |
| COPP does not cause CPI | 390.212 | 0.000 | COPP→CPI | 21.470* | 0.000 | COPP→CPI |
| WTI does not cause CPI | 64.963* | 0.000 | WTI→CPI | 19.211* | 0.000 | WTI→CPI |
| COPP does not cause WTI | 34.458* | 0.000 | COPP→WTI | 25.446* | 0.000 | COPP→WTI |
| CPI does not cause WTI | 14.382* | 0.000 | CPI→WTI | 17.271* | 0.000 | CPI→WTI |
| Break Date: 2003-2008 (except) | | 1991-2003 | | | 2008-2020 | |
| CPI does not cause COPP | 3.1236 | 0.373 | - | 24.427* | 0.000 | CPI→COPP |
| WTI does not cause COPP | 4.3108 | 0.230 | - | 13.262* | 0.004 | WTI→COPP |
| COPP does not cause CPI | 2980.00 | 0.000 | COPP→CPI | 27.390* | 0.000 | COPP→CPI |
| WTI does not cause CPI | 1294.3* | 0.000 | WTI→CPI | 25.576* | 0.000 | WTI→CPI |
| COPP does not cause WTI | 32.186* | 0.000 | COPP→WTI | 50.706* | 0.000 | COPP→WTI |
| CPI does not cause WTI | 31.382* | 0.000 | CPI→WTI | 56.126* | 0.000 | CPI→WTI |
| | | | | 24.427* | 0.000 | CPI→COPP |

*Statistically significant at 5% level

range to catch clearly structural changing in model parameters. It was seen slow changing in parameters in range. The results show there are obviously different constant and slope parameters between pre-2003 and post-2008 samples. In cereal price equation, the constant parameter has decreased by approximately 50% in post-2008 period. This means that average price performers of food price, now more related to other commodity prices in the system, not based itself as before. In post-2008 copper price

equation, we found that only small lagged copper price has positive effect on cereal price according to pre-2003 sample.

In pre-2003 period, results shows that there is uni-directional causality relation from copper to cereal and WTI crude oil, bi-directional relation between cereal and WTI crude oil. In post-2008 period, all commodity trio prices series have bi-directional causality relation among each two sub-blocks, energy-food,

materials-food and energy-materials. We found that parameters and causal relations changed in post-2008. Now, there is a relation running from cereal and WTI crude oil to copper price, bi-directionally. While, in pre-break range, WTI and cereal have in-directional effect on copper, in post-break range sample, directional. And, we found, the changes of series causality relation realized in step by step in breaks range of 2003-2008.

5. CONCLUDING REMARKS

The study aims to investigate time series properties and causal relation among commodity trios prices for the period of 1991-2020. We generate that sub-break samples using by structural break test and try to catch break point of series and causal equation parameters. For a more detailed analysis and capture clues, it would be usefull to select the best proxy from all contents of commodity variables. We generated three proxy variable instead of materials, energy and food prices using factor analysis; as copper, WTI crude oil and cereal price.

Our empirical findings can be summarized as follows. Unit-root test results show that commodity price series are stationary with break. BDS linearity/non-linearity test presents all series are linear in sub-break samples. Bai and Perron (2003) test result exhibits materials, energy and food price series has one important break in 2003, 2004 and 2005, respectively, due to 9/11 and 2003 Gulf War shocks and lagged effect. Johansen's trace test suggest the presence of a long-run relationship among the energy, food, materials prices, existence of cointegration. Finally, we have estimated vector autoregression model taking into account moving break period by 2003-2008. We found that model parameters exhibited changing process with 2003-2008 break range. In pre-2003 period, results shows that there is uni-directional causality relation from copper (materials) to cereal (food) and WTI crude oil (energy), bi-directional relation between cereal (food) and WTI (energy). In post-2008 period, all commodity trio prices series have bi-directional causality relation between all two sub-blocks; energy-food, materials-food and energy-materials. We found that series and causal relations have been changed in post 2003-2008 break range. Now, there is a relation running from cereal (food) and WTI (energy) to copper (materials) price, bi-directionally. While, in pre-break, WTI and cereal have in-directional effect on copper, in post break sample, directional. And, we found, the changes of series causality relation realized in step by step in breaks range of 2003-2008.

From a philosophical point of view, there are two possible options in the face-of crisis/break-producing economic system; to go on to fight inside facit circle or break out. Inside circle, risk management process requires more complex analysis system. Break out's one is related with learning-unstucking ability and evolution process. We need creating an "Unlife Life" quadrant perspective: Learn/Unlearn-Stuck/Unstuck boxes. We need to know what should we know and how can be release old systematics. And need more progress in world internatinal relation to pass advance and

align box/stages. Sustainable source come to the agenda instead of utopic's one "equitable sharing of natural source" Creating unlife-life path for the world economy means that sustainable game. Sustainability is important in materials and food as well as energy source.

This approach can provide further insights into the dynamics of commodity prices and their interdependencies. We show that commodity prices faced to break/s based on the results of mostly source rich countries' excessive interventions in market. Now, commodity prices have moved the new atmosphere, the more under directly effect eachother.

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